Triggering of Myocardial Infarction by Cocaine

Murray A. Mittleman, MD, DrPH; David Mintzer; Malcolm Maclure, ScD; Geoffrey H. Tofler, MB; Jane B. Sherwood, RN; James E. Muller, MD

Background—Cocaine has been implicated as a trigger of acute myocardial infarction in patients with and those without underlying coronary atherosclerosis. However, the magnitude of the increase in risk of acute myocardial infarction immediately after cocaine use remains unknown.

Methods and Results—In the Determinants of Myocardial Infarction Onset Study, we interviewed 3946 patients (1282 women) with acute myocardial infarction an average of 4 days after infarction onset. Data were collected on the use of cocaine and other potential triggers of myocardial infarction. We compared the reported use of cocaine in the hour preceding the onset of myocardial infarction symptoms with its expected frequency by using self-matched control data based on the case-crossover study design. Of the 3946 patients interviewed, 38 (1%) reported cocaine use in the prior year and 9 reported use within the 60 minutes preceding the onset of infarction symptoms. Compared with nonusers, cocaine users were more likely to be male (87% vs 67%, \(P=0.01\)), current cigarette smokers (84% vs 32%, \(P<0.001\)), younger (44±8 vs 61±13 years, \(P<0.001\)), and minority group members (63% vs 11%, \(P<0.001\)). The risk of myocardial infarction onset was elevated 23.7 times over baseline (95% CI 8.5 to 66.3) in the 60 minutes after cocaine use. The elevated risk rapidly decreased thereafter.

Conclusions—Cocaine use is associated with a large abrupt and transient increase in the risk of acute myocardial infarction in patients who are otherwise at relatively low risk. This finding suggests that studying the pathophysiological changes produced by cocaine may provide insights into the mechanisms by which myocardial infarction is triggered by other stressors. (Circulation. 1999;99:2737-2741.)

Key Words: cocaine • myocardial infarction • epidemiology

Cocaine has been implicated as a trigger of acute myocardial infarction both in patients free of coronary artery disease1 and more frequently in patients with underlying coronary atherosclerosis.1–4 Despite these reports, there are no published data from controlled studies evaluating the magnitude of the increase in risk of acute myocardial infarction immediately after cocaine use. With more than 30 million Americans having tried cocaine at least once and a reported 5 million regular users,5 a better understanding of the magnitude of this cardiac risk is important from a public health perspective.

Although the mechanisms involved in cocaine-induced myocardial infarction are not well understood, it is known that cocaine blocks the presynaptic reuptake of norepinephrine and dopamine, leading to high concentrations of these neurotransmitters at postsynaptic receptor sites.6 This adrenergic stimulation increases myocardial oxygen demand by causing an increase in heart rate, blood pressure, and left ventricular contractility.7 In addition, acute exposure to cocaine has been documented to cause coronary vasoconstriction,8–16 which, together with marked increases in arterial pressure, may lead to disruption of atherosclerotic plaques in the coronary vasculature that are vulnerable to the increase in shear forces.17,18 Furthermore, acute administration of cocaine has been reported to increase platelet aggregability19,20 and may lead to in situ thrombus formation.1,2,4,21–24

To evaluate the magnitude of the risk of having an acute myocardial infarction triggered by the recreational use of cocaine, we collected data on the use of cocaine in 3946 patients (1282 women) with acute myocardial infarction interviewed for the Determinants of Myocardial Infarction Onset Study. In this multicenter, interview-based study, we compared the reported use of cocaine in the hour preceding the onset of myocardial infarction symptoms with its expected frequency by using self-matched control data based on the case-crossover study design.

Methods

Study Population

Between August 1989 and September 1996, a total of 3946 patients (2664 men and 1282 women, age range 20 to 92 years) were
interviewed at 64 medical centers a median of 4 days after myocardial infarction.

Interviewers identified eligible cases by reviewing coronary care unit admission logs and patients’ charts. For inclusion in the study, patients were required to meet all of the following criteria: at least 1 creatine kinase level above the upper limit of normal for the clinical laboratory performing the test, positive MB isoenzymes, an identifiable onset of pain or other symptoms typical of infarction, and the ability to complete a structured interview. The protocol was approved by the institutional review board at each participating center, and informed consent was obtained from each patient.

Detailed chart reviews and patient interviews were conducted by trained research personnel, as previously described. Data were collected on standard demographic variables as well as risk factors for coronary artery disease. The interview identified the time, place, and quality of myocardial infarction pain and other symptoms as well as the timing and estimated usual frequency of exposure to potential triggers of myocardial infarction onset during the prior year. In addition, patients were asked if they had used cocaine in the year preceding their infarction. Patients who reported any use of cocaine were also asked to report the last time that they had used cocaine and their usual frequency of using cocaine over the prior year.

**Study Design**

The design of the Onset Study has been described in detail elsewhere. In brief, we used a case-crossover study design to assess the change in risk of acute myocardial infarction during a brief “hazard period” after exposure to cocaine and other potential triggers of myocardial infarction onset. An important feature of the case-crossover design is that control information for each patient is based on his or her own past exposure experience. Self-matching results in freedom from confounding by risk factors that are stable over time but often differ between study subjects.

Cocaine use in the hazard period, the 1-hour period immediately preceding the onset of myocardial infarction symptoms, was compared with its expected frequency based on control data obtained from the patients. We used the usual frequency of cocaine use over the year prior to myocardial infarction to estimate its expected frequency in an average 1-hour period in this patient population.

**Statistical Analysis**

The analysis of case-crossover data are an application of standard methods for stratified data analysis. In this analysis, the stratifying variable is the individual patient, as in a crossover experiment. The ratio of the observed exposure frequency in the hazard period to the expected frequency (from the control information) was used to calculate estimates of the odds ratio as a measure of relative risk. The amount of person-time exposed to cocaine was estimated by multiplying the reported usual annual frequency of exposure by the duration of its hypothesized physiological effect (1 hour). Unexposed person-time was then calculated by subtracting the exposed person-time in hours from the number of hours in 1 year. The data were analyzed with the use of methods for cohort studies with sparse data in each stratum.

**Results**

The characteristics of the patients interviewed are presented in Table 1. Of the 3946 patients with myocardial infarction who were interviewed, 38 (1%) reported that they had used cocaine in the year preceding their myocardial infarction. Compared with nonusers, cocaine users were more likely to be male (87% vs 67%, P = 0.01), current smokers (84% vs 32%, P < 0.001), younger (44 ± 8 vs 61 ± 13 years, P < 0.001), and members of a minority group (63% vs 11%, P < 0.001).

Table 2 shows the distribution of the usual frequency of cocaine use among the 38 patients who reported using cocaine in the year before their myocardial infarction. Of the 38 patients, 9 reported using cocaine within the 60 minutes before the onset of their infarction symptoms. In addition to these 9 patients, 1 patient reported using cocaine between 60 and 120 minutes before the onset of symptoms, and 1 additional patient reported cocaine use in the period from 120 to 180 minutes before myocardial infarction onset.

Within 1 hour after using cocaine, the risk of myocardial infarction onset was elevated 23.7 times (95% CI 8.5 to 66.3). The Figure shows that the relative risk of myocardial infarction was much lower during the second and third hours after cocaine use (controlling for subsequent exposure). These relative risks were not statistically significantly elevated. However, the confidence intervals for these relative risks are wide, and persistence of a moderately increased risk cannot be ruled out on the basis of these data.

**Discussion**

These results from the Onset Study corroborate the reports from case-series that cocaine use can trigger the onset of myocardial infarction. The conclusions of the Onset Study differ from those of previous studies of case-series that cocaine use can trigger the onset of myocardial infarction.
by guest on April 20, 2017 http://circ.ahajournals.org/ Downloaded from

and dopamine, leading to a high concentration of these neurotransmitters at post synaptic receptor sites. This adrenergic stimulation in turn causes an increase in heart rate, blood pressure, and left ventricular contractility. Second, within 15 minutes of intranasal administration of even low doses of cocaine, coronary vasoconstriction occurs in both stenotic and nonstenotic segments through stimulation of \( \alpha \)-adrenergic receptors. In addition, cocaine has been shown to increase the endothelial release of endothelin, probably through stimulation of \( \alpha \)-receptors. Animal models indicate that cocaine administration leads to a rapid rise in intracellular free Ca\(^{2+}\) and a concomitant loss of intracellular Mg\(^{2+}\) in vascular smooth muscle cells. These fluxes in divalent cations may directly contribute to vasoconstriction. Third, cocaine has been documented to cause an increase in platelet aggregability in vivo and in vitro testing. Furthermore, angiographic studies have shown that some patients who had myocardial infarction after cocaine use had occlusive thrombi at nonstenotic sites within their coronary arteries. Finally, accelerated atherosclerosis has been detected in young cocaine users. Although speculative, it is possible that the development of such subclinical disease may contribute to the likelihood that a habitual cocaine user will have vulnerable atherosclerotic plaques present in their coronary vessels at the time of subsequent cocaine use.

In addition to its effects on coronary arteries, cocaine can cause an acute deterioration in left ventricular systolic and diastolic function. Pitts et al recently demonstrated that an intracoronary infusion of cocaine can lead to an increase in left ventricular end-diastolic pressure and left ventricular end-systolic volume as well as a decrease in left ventricular ejection fraction. A possible mechanism for these effects is that cocaine may alter Ca\(^{2+}\) handling by myocytes. For example, Perreault et al demonstrated that cocaine can cause an increase in cytosolic Ca\(^{2+}\) concentration in isolated myocardial tissue. This in turn leads to prolongation of the calcium current and a negative lusitropic effect.

The present study has several potential limitations. No data were collected on the method of cocaine exposure, for example, intravenous, intranasal, or inhaled crack cocaine use. In addition, because the data are based on patient self-report, some misclassification of exposure probably occurred. For example, patients may be reluctant to report that they had used cocaine before their myocardial infarction. The effect of such a bias would be to reduce the magnitude of the estimated relative risk. In an effort to minimize such reporting bias as well as to maintain patient confidentiality, efforts were made to ensure the patient’s privacy during the interview. Furthermore, to obtain comparable reporting of cocaine use for all of the hourly intervals during the day preceding the infarction, patients were not informed of the duration of the hypothesized hazard period.

There is a possibility of bias caused by differential survival of cases who had a myocardial infarction triggered by different mechanisms. For example, if patients whose infarctions were triggered by cocaine were more likely to die than those whose infarctions were unrelated to cocaine, then the apparent relative risk may be underestimated.

On the other hand, it is possible that patients underreport the use of cocaine because of the social stigma attached to it, but they may be relatively accurate about its use on the day of their myocardial infarction because of potential clinical benefit. This may result in an overestimate of the relative risk. One way to control for such recall bias is to use the period from 1 to 3 hours before the infarction as the control period. In this case the null hypothesis is that the 11 episodes of cocaine use would be evenly distributed over the 3 hours before myocardial infarction onset after cocaine use. Relative risk of myocardial infarction onset plotted on a logarithmic scale for each of the 3 hours after cocaine use. Relative risks were estimated by comparing the frequency of cocaine use in each of the 3 hours before myocardial infarction onset to its expected frequency based on each patient’s reported usual frequency of cocaine use over the prior year. Error bars indicate 95% CI. Dotted line represents baseline risk during periods of nonexposure to cocaine.

Relative risk of myocardial infarction onset after cocaine use. Relative risk of myocardial infarction onset is plotted on a logarithmic scale for each of the 3 hours after cocaine use. Relative risks were estimated by comparing the frequency of cocaine use in each of the 3 hours before myocardial infarction onset to its expected frequency based on each patient’s reported usual frequency of cocaine use over the prior year. Error bars indicate 95% CI. Dotted line represents baseline risk during periods of nonexposure to cocaine.
before the onset of myocardial infarction symptoms. What we observed was 9 exposures in the first hour and 1 episode per hour in each of the other 2 hours. This corresponds to a relative risk of 9.0 (95% CI 1.9 to 41.7; P<0.001).

Because of the small number of exposed cases, we were unable to evaluate whether the risk of having a myocardial infarction differed in subsets of patients. For example, on the basis of the work of Moliterno et al, the combination of cocaine use and cigarette smoking might be particularly harmful. Similarly, we were unable to determine whether the risk of sustaining a cocaine-associated myocardial infarction differed for frequent versus infrequent users.

In traditional epidemiological studies of coronary heart disease, confounding by differences in risk factors between individuals is a major threat to validity. A strength of the case-crossover design used in this study is that self-matching ensures that within strata there is no variability in traditional chronic risk factors for coronary heart disease. Thus, by design, confounding by all traditional chronic risk factors for coronary heart disease, whether measured or unmeasured, is controlled for in the analysis.7,37,31

A limitation of the case-crossover design used in this study is that like case-control studies, the absolute risk of myocardial infarction onset cannot be directly estimated from the data. However, an estimate of the baseline risk can be made with the use of other data sources. For example, on the basis of the Framingham Heart Study risk equation,46,47 the baseline risk of acute myocardial infarction for a typical cocaine user in this study (44-year-old male smoker with average risk of acute myocardial infarction for a typical cocaine user in this study (44-year-old male smoker with average risk of acute myocardial infarction) differed for frequent versus infrequent users.

Acknowledgments

This work was supported by grants from the National Heart, Lung, and Blood Institute (HL-41016) and the American Heart Association (9630115N). We are grateful to the Onset Study interviewers for their dedication and Diane Walkoff, BS, for expert assistance in conducting the Onset Study.

References


